



**Placebo effects in hearing aid trials are reliable**

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Manuscripts

Dawes *Reliability of placebo effects*

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1 **Placebo effects in hearing aid trials are reliable**

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10 Key words: placebo effect, reliability, hearing aid trial

11 Abbreviations: FAAF, Four Alternative Auditory Feature test, NAL-NL1; National Acoustics

12 Laboratories-Non-Linear 1

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**Abstract**

Objective: A recent study suggested that placebo effects are a source of bias in non-blinded hearing aid trials. Given the potential impact of this finding on the interpretation of non-blinded trials and design of future research trials, the objective of the present study was to investigate the reliability of this effect.

Design: Using the same procedure as an earlier study, participants were told that they were taking part in a trial of new hearing aid technology. Participants compared two devices that were acoustically identical, except one was described as “new” and the other as “conventional”. Participants completed a speech-in-noise test, sound quality ratings and rated overall personal preference for both hearing aids.

Study sample: Sixteen adult hearing aid users.

Results: Participants had significantly better mean speech-in-noise performance (70.9% versus 66.8%,  $Z=2.30$ ,  $p=0.02$ , effect size Pearson’s  $r=0.15$ ) and sound quality ratings for the “new” hearing aid (8.1 versus 7.4,  $Z=-2.99$ ,  $p=0.003$ ,  $r=0.28$ ). A significant proportion of participants (75%) expressed an overall preference for the “new” hearing aid ( $p=0.001$ , effect size  $\phi_c=0.66$ ).

Conclusion: Placebo effects reliably impact on hearing aid trials. In order to control for placebo effects, double-blind methodology is optimal. However, when double-blinding is not possible other strategies may be appropriate.

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3 39 Placebo effects are clinical responses associated with the expectation surrounding  
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5 40 treatment, rather than with any intrinsic property of the treatment. Placebo effects can  
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7 41 occur in conjunction with active treatments, and either facilitate or inhibit them (Luparello  
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9 42 et al., 1968; Colloca et al., 2004; Benedetti et al., 2007). In relation to clinical medicine, it is  
10  
11 43 suggested that placebo effects could be (and are) used to optimise outcomes for patients.  
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13 44 Such positive effects might be encouraged by the use of suggestion or through positive  
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15 45 interaction between the clinician and the patient (Thompson, 2000; Benedetti, 2011; Finniss  
16  
17 46 et al., 2011). Placebo effects are also widely recognised in medical research, and control  
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19 47 conditions with double-blind methodology<sup>1</sup> are routinely incorporated into clinical trials  
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21 48 (Thompson, 2000; Price et al., 2008). However, in clinical audiology and audiological  
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23 49 research, placebo effects are not typically considered. There is some research evidence to  
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25 50 suggest that they may have relevance for audiology, as follows.  
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31  
32 51 Around the time that digital hearing aids were being first introduced, Bentler and colleagues  
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34 52 (2003), investigated whether labelling hearing aids as 'digital' impacted on measures of  
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36 53 hearing aid benefit and overall preference. They provided participants with two identical  
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38 54 hearing aids that were labelled as 'digital' (versus 'conventional'). Participants used each  
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40 55 hearing aid for one month. They then completed a battery of hearing aid benefit measures,  
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42 56 involving speech perception and self-report measures. Effects were generally small; the  
43  
44 57 effect of labelling accounted for 2 to 32 % of variance of the outcome measures. Differences  
45  
46 58 between labelling conditions were not statistically significant individually (with the  
47  
48 59 exception of some subscales from the Abbreviated Profile of Hearing Aid Benefit; Cox &  
49  
50 60 Alexander, 1995). Overall, a statistically significant majority of participants (33 of 40)

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<sup>1</sup> Double-blind studies are those where neither the participant nor the experimenter is aware which the experimental condition is and thus control for the possible affect of expectation.

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3 61 reported that they preferred the 'digital' hearing aid. Bentler et al concluded that the  
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5 62 participants' expectation that 'digital' hearing aids must be better had affected both  
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8 63 performance of measures of hearing aid benefit and their overall preference.  
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11 64 Dawes et al (2011) went on to investigate the impact of patient expectation on the outcome  
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13 65 of a trial of two hearing aids. The aim of Dawes et al's study was to test whether the effect  
14  
15 66 of expectation observed by Bentler and colleagues was specific to the 'digital' label, or  
16  
17 67 whether a more general label of 'new technology' could set up positive expectations about  
18  
19 68 performance, with a consequent impact on outcome measures. If this were the case, it  
20  
21 69 would be relevant for any trial of new hearing aid technology. Twenty experienced adult  
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23 70 hearing aid users were told that they were taking part in a trial of new hearing aid  
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25 71 technology. In a single test session, users completed a speech-in-noise test, sound quality  
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27 72 ratings and overall personal preference. Two hearing aids were compared, one being  
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29 73 described as 'new' and the other 'conventional'. In reality, both hearing aids were identical  
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31 74 and were programmed to the same prescription target for a typical age-related hearing loss.  
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33 75 SIN performance was marginally better (by around 2%) for the 'new' hearing aid for the  
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35 76 speech-in-noise test, sound quality ratings were statistically significantly higher for the 'new'  
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37 77 hearing aid, and 15 out of 20 participants expressed an overall personal preference for the  
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39 78 'new' hearing aid (with the remainder expressing no preference). The interpretation was  
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41 79 that describing one hearing aid as being 'new' had set up an expectation in the minds of the  
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43 80 participants that performance would be better for the 'new' hearing aid, and that  
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45 81 participants' expectation had impacted upon performance of the test measures. On the  
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47 82 basis of Bentler et al's (2003) study and Dawes et al's (2011) findings, Dawes et al suggested  
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3 83 a need to control for placebo effects and advised caution in interpreting the results of trials  
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5 84 that did not control for placebo effects.  
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9 85 One area of uncertainty is whether expectation does have a reliable impact on performance  
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11 86 and measures of hearing aid benefit. In Bentler et al's study, effects on individual tests (with  
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13 87 the exception of some subscales for the self-report questionnaire) were generally small and  
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15 88 statistically non-significant (although there was a significant effect overall). In Dawes et al's  
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17 89 study, the difference in performance on the speech in noise test was small and statistically  
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20 90 non-significant (although the authors argued that this small difference was significant in the  
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22 91 context of positive findings for the other measures and in hearing aid trials generally, where  
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24  
25 92 small effects are typically sought).  
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28 93 If placebo effects are reliable in the context of aided performance and measures of hearing  
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30 94 aid benefit, there would be an obvious need to account for placebo effects in the design of  
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33 95 future hearing aid trials. Such effects may also have relevance for clinical audiology, if for  
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35 96 example participant expectation might be utilised to increase the benefit of audiological  
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37 97 treatments. Thus, the objective of the present study was to investigate the reliability of the  
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39 98 effect observed by Dawes et al (2011) by replicating that study in a different group of  
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41 99 participants and with a different experimenter administering the test measures. The  
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44 100 hypothesis was that if placebo effects on hearing aid trials are reliable, a similar pattern of  
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46 101 results to that observed by Dawes et al (2011) would be obtained. This would add weight to  
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48 102 the suggestion that placebo effects should be controlled for in trials of hearing aid (or  
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51 103 similar) technology.  
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55 104 **Method**  
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3 105 Methodology follows that used by Dawes et al (2011) but is summarised below.  
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6 106 *Participants*  
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9 107 A sample size was chosen to be comparable with Dawes et al's (2011) study and was  
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11 108 intended to be similar to a typical of hearing aid trial, which use small to medium sample  
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13 109 sizes that are sufficient to detect effects that are large enough to be of clinical relevance.  
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15 110 Accordingly, sixteen participants aged between 61 to 86 yr (M = 76 yr, SD = 7 yr) were  
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17 111 recruited from a local hospital-based audiology clinic. Inclusion criteria were: i) at least 12  
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19 112 months daily hearing aid use, ii) English as a first language, iii) symmetrical, mild-to-  
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21 113 moderate, sloping high frequency sensorineural hearing loss of at least 45 dB HL at 2 – 6  
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23 114 kHz, iv)  $\leq 5$  dB difference between the ears at two or more adjacent frequencies between  
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25 115 0.25 and 8kHz and v) normal middle ear function. Pure-tone audiometry was performed in a  
26  
27 116 sound-treated room using a calibrated Kamplex KLD 21 audiometer with TDH-39  
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29 117 headphones and a B-71 bone vibrator. Normal middle ear function was confirmed with a GSI  
30  
31 118 38 Auto Tym. All participants were unilateral hearing aid users. The reason for this was that  
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33 119 unilateral hearing aids are typically prescribed by the audiology clinic from which  
34  
35 120 participants were recruited. When invited to take part, participants were told that the  
36  
37 121 purpose of the study was to evaluate new hearing aid technology. On completion of the  
38  
39 122 study, participants were informed about the true purpose of the study. Ethical approval was  
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41 123 obtained from the NHS Central Manchester Research Ethics Committee. Written consent  
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43 124 was obtained from all participants. Participants were reimbursed travel costs but no other  
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45 125 compensation was provided.  
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127 *Test hearing aids*

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3 128 Test hearing aids were two Starkey A312 Strata behind-the-ear digital aids with seven band,  
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5 129 three channel wide dynamic range compression with a noise reduction algorithm. One  
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8 130 hearing aid was to be referred to as the 'new' one while the other was referred to as the  
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10 131 'conventional' aid. The 'new' hearing aid had a yellow case while the 'conventional' aid had  
11  
12 132 a beige case. The hearing aids were programmed to the same NAL-NL1 prescription target  
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14 133 based on an audiogram for a typical age-related hearing loss (35 dB HL at 500Hz, 40 dB HL at  
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16 134 1000Hz, 50 dB HL at 2000 Hz, 60 dB HL at 4 kHz, and 80 dB HL at 8 kHz). Casings were  
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19 135 switched between hearing aids so that for half the group, one hearing aid had a yellow case  
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21 136 and was the 'new' hearing aid, while the other had a beige case and was the 'conventional'  
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23 137 aid. For the other half of the group, casings and designation (new/conventional) was  
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26 138 switched. This procedure was used to control for any actual acoustic differences between  
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29 139 hearing aids.

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33 141 In order to ensure that the two hearing aids produced identical amounts of gain, repeated  
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35 142 coupler measurements were obtained after initial programming, after switching cases and  
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37 143 after completion of testing. All measurements for the two hearing aids at all frequencies  
38  
39 144 were consistently within 1dB. Coupler gain targets and repeated gain measurements are  
40  
41 145 shown in Table 1. Listening tests confirmed the similarity of the hearing aids. All participants  
42  
43 146 used their current ear mould (all hard acrylic with a pressure equalization vent) with the test  
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45 147 hearing aids. Test hearing aids were thus trialled with the ear that was normally fitted with  
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48 148 the hearing aid.

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51 149 (Table 1 here)

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53 150 *Outcome measures*  
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Dawes *Reliability of placebo effects*

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3 151 Participants were seated in a comfortable chair 1.5m from a loudspeaker at 0 degrees  
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5 152 azimuth. Sound levels for each stimulus refer to the level measured at the reference point,  
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7 153 defined as the centre of the participants head with the participant absent.  
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13 155 *Speech in noise test*

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15 156 The Four Alternative Auditory Feature test (FAAF; Foster and Haggard, 1987) is a  
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17 157 computerised 80 item single-syllable, closed set word recognition test. Participants hear the  
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19 158 sentence 'Can you hear X clearly?' and are required to choose from a selection of four  
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21 159 words on a screen using a mouse click to identify the word that they heard. The test was  
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23 160 administered with the sentences at 65 dB (A) in the presence of speech-shaped noise at +2  
24  
25 161 dB SNR. A practice list of 12 words was used to familiarise participants with the test  
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27 162 procedure with participants using their own hearing aid. Participants completed separate  
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29 163 runs of the FAAF with both the new and the conventional aid, with order of testing counter-  
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31 164 balanced across participants.  
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38 166 *Sound quality rating test*

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40 167 Participants listened to six sound samples and rated them on clarity, comfort and overall  
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42 168 impression using a 10 point visual analogue scale (based on Arlinger et al, 1998). Sound  
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44 169 samples were Bamford-Kowal-Bench sentences (BKB; Bench et al, 1979) spoken by male and  
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46 170 female voices in quiet and in noise, music and an environmental sound (robin song). All  
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48 171 samples were 10 seconds long, digitized at 44 kHz. Samples were equalised to have the  
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50 172 same long-term RMS power and presented at 65 dB (A). BKB sentences in noise were  
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52 173 presented at +2 dB SNR in broadband noise at the same overall presentation level for the  
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54 174 sentences in isolation (65 dBA).  
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3 1754  
5 176 *Personal preference*6  
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8 177 At the end of the test session, participants were asked to indicate if they had an overall9  
10 178 personal preference for either hearing aid by choosing one of three categories: 'the new11  
12 179 hearing aid is best', 'the conventional hearing aid is best' or 'I cannot tell any difference'.13  
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15 18016  
17 181 *Procedure*18  
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20 182 At the beginning of the session, participants were given an explanation of the (false) aim of21  
22 183 the study, i.e. to evaluate new hearing aid technology and were shown the two hearing aids.23  
24 184 The hearing aid with a yellow case was introduced as the one containing 'new technology',25  
26 185 while the hearing aid with a beige case was introduced as the 'conventional' hearing aid. No27  
28 186 negative comments were made about the 'conventional' hearing aid. Otoscopy,29  
30 187 tympanometry and pure tone audiometry was carried out first, followed by FAAF test,31  
32 188 sound quality ratings and personal preference. The order of testing was counterbalanced so33  
34 189 that half the participants performed the tasks with the 'new' hearing aid first, while the35  
36 190 other half performed them with the 'conventional' aid first. All measurements were made37  
38 191 within a single test session of around one hour's duration.39  
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44 192 *Statistical analysis*45  
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47 193 Non parametric tests (Wilcoxon signed-rank test) were selected for FAAF test because of48  
49 194 non-normal distribution of data (positive skew) and to allow comparability with Dawes et al50  
51 195 (2011). Performance was above chance levels for all participants, and because relative52  
53 196 differences between two conditions are of interest rather than absolute performance, data54  
55 197 from all participants was retained and appropriate non-parametric tests applied. Non-56  
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198 parametric tests were also applied for sound quality ratings, as appropriate for ordinal data.  
199 To test for an effect of the order of testing (i.e. new or conventional first), Mann-Whitney U  
200 test was applied for both FAAF and sound quality ratings. A Chi-squared test was used to  
201 test whether the proportion who reported a preference for either hearing aid was different  
202 to that expected by chance. In order to estimate the magnitude of the effect of expectation  
203 and to facilitate comparison with effects observed by Dawes et al (2011) and with  
204 experimental effects sought in clinical hearing aid trials, effect sizes were calculated using  
205 Pearson's  $r$  statistic<sup>2</sup> using the means and standard deviations for the respective variables.  
206 Cramér's  $V^3$  (denoted as  $\phi_c$ ) was calculated for the Chi-squared test.

## 208 **Results**

209 There were no significant effects of test order (new or conventional first) on FAAF score or  
210 sound quality ratings;  $p$ 's 0.21 to 0.87.

### 211 *FAAF test*

212 Table 2 shows mean performance for the 'new' and the 'conventional' hearing aids in the  
213 present study and in Dawes et al (2011). In the current study, performance was statistically  
214 significantly higher for the 'new' hearing aid (compared to the 'conventional' hearing aid,  
215 with the size of the difference approximately double that reported in Dawes et al (2011). In  
216 the current study, twelve participants (75%) demonstrated a better performance in 'new'  
217 hearing aid condition compared to the 'conventional' hearing aid condition. One participant

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<sup>2</sup> The absolute magnitude of Pearson's  $r$  varies between 0 and 1, with 1 indicating a perfect relation between the two variables and 0 indicating no relation. As a guide, effects of 0.2 are considered 'small', 0.5 'medium' and 0.8 'large'.

<sup>3</sup> Cramér's  $V$  is applied to goodness-of-fit chi-squared models when there is a 1xk table (i.e. degrees of freedom are greater than 1). Like the Pearson  $r$  statistic, it provides a measure of association between two variables that varies in size between 0 and 1.

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3 218 (6%) performed equally in both conditions and three participants (19%) performed worse in  
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5 219 the 'new' condition than the 'conventional' condition.

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8 220 (Table 2 here)

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11 221 *Sound quality rating test*

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13 222 Median scores for each of the six different sound samples played to the participants for  
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15 223 each dimension of clarity, comfort and overall impression were obtained for each  
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17 224 participant. Both hearing aids scored well, although the 'new' hearing aid scored statistically  
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19 225 significantly higher than the 'conventional' hearing aid. Overall sound quality rating (median  
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21 226 of all subjective ratings) was also significantly higher for the 'new' hearing aid compared to  
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23 227 the 'conventional' hearing aid and this was statistically significant. Table 3 details mean  
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25 228 ratings for each dimension and results of statistical comparisons for the present study and  
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27 229 for Dawes et al (2011). The pattern of better ratings for the 'new' hearing aid is common to  
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29 230 both studies.

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31 231 (Table 3 here)

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33 232 *Overall preference*

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35 233 Twelve participants stated a preference for the 'new' hearing aid, while four reported that  
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37 234 they could not tell the difference between the two hearing aids. None preferred the  
38  
39 235 'conventional' hearing aid. This difference is statistically significant ( $\chi^2(2) = 14.0, p = 0.001,$   
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41 236  $\phi_c = 0.66$ ). The proportion of those who preferred the 'new' hearing aid is the same as that  
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43 237 reported by Dawes et al (2011); 75%.

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46  
47 239 **Discussion**

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3 240 Overall, a similar pattern of findings was obtained in the current study compared to Dawes  
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5 241 et al (2011); participants performed better on the speech in noise test and made more  
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7 242 favourable sound quality ratings for the 'new' hearing aid than the 'conventional' one. A  
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9 243 significantly higher proportion of participants preferred the 'new' hearing aid. Effect sizes  
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11 244 were larger for sound quality ratings and overall preference (measures that rely on  
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13 245 participant self-report) over the speech in noise test. As with Dawes et al (2011), the  
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15 246 interpretation of these findings is that describing one hearing aid as 'new' set up  
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17 247 expectations in the participants that this hearing aid must perform better than the  
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19 248 'conventional' one, and that this expectation impacted upon participant's performance of  
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21 249 the experimental tasks. Such an expectation may have resulted in improved performance  
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23 250 with the 'new' hearing aid or worse performance with the 'conventional' one (consistent  
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25 251 with a 'nocebo' effect; Benedetti et al, 2007) or both. Thus, participant expectation appears  
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27 252 to have a reliable effect on performance.  
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34 253 In the current study, FAAF scores tended to be higher while sound quality ratings tended to  
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36 254 be lower than in Dawes et al's (2011) study. This difference is unlikely to be due to  
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38 255 difference in procedure; hearing aids, tests and procedure were identical. The primary  
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40 256 differences between the two studies are i) a different group of participants and ii) a  
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42 257 different experimenter. It seems unlikely that the influence of the experimenter could be a  
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44 258 primary reason for this difference. If there was a systematic effect of the experimenter, the  
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46 259 direction of the effect should be the same for both measures. This difference between  
47  
48 260 studies is more likely to be due to uncontrolled differences between the participants in the  
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50 261 two studies, such as the average level of hearing loss.  
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3 262 As with Dawes et al (2011), because of time and ethical constraints, outcome measures  
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5 263 were only made within a single test session in the current study. The duration of effects  
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7 264 seen in current study and in Dawes et al (2011) are therefore not known. Placebo effects in  
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10 265 clinical medicine can be very long lasting (effects of up to a year have been recorded;  
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12 266 Dimmons et al. 1960). Bentler et al (2003) found that the influence of labelling a hearing aid  
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14 267 as 'digital' persisted after one month of hearing aid use, and so it is possible that the effects  
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17 268 of expectation and labelling on hearing aid benefit may be long lasting. This is a topic for  
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19 269 future research (see also the section on the role of placebo effects in clinical audiology  
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21  
22 270 below).

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25 271 If placebo effects do have a reliable impact on trials of hearing aids (or similar technology),  
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27 272 then trials should include controls for placebo effects. This is necessary so that one may  
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29 273 have confidence that any benefit associated with the experimental condition is not at least  
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31 274 partly due to a placebo effect. But how should one control for placebo effects in hearing aid  
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33 275 trials? A randomised controlled trial is the gold standard methodology for clinical trials  
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36 276 (Benedetti, 2009). In a randomised controlled trial, participants are randomly allocated to  
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38 277 test and control conditions. However, this does not eliminate placebo effects because if  
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40 278 participants or experimenters are aware of the identity of the groups, there is the potential  
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43 279 for the expectation of the participant and/or the experimenter to impact on the outcome  
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46 280 (Gracely et al, 1985). The gold standard control for placebo effects is a double-blinded  
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48 281 design in which neither participants nor experimenters are aware of the identity of the  
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50 282 control or experimental condition. A double-blinded design means that neither the  
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52 283 participant's nor the experimenter's expectations should influence the outcome of the  
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55 284 study. This design may be applicable for many instances in audiological research. For  
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3 285 example, two signal processing schema could be compared in a double-blinded design using  
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5 286 identical hardware such as a programmable digital hearing aid or via headphone simulation.  
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7 287 Researchers wishing to apply the most rigorous methodology should consider trials of new  
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10 288 technology that would be amenable to double-blind designs.  
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13 289 However, an experimental hearing aid may look physically different to the comparison one,  
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15 290 in which case a double-blind design may not be possible. In this case, it would be desirable  
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17 291 to establish an alternative to double-blind methodology. One method of minimising placebo  
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19 292 effects may be employing an experimenter who would provoke minimal labelling effects.  
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22 293 Clinician characteristics such as warmth, empathy, prestige and friendliness have been  
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24 294 demonstrated to impact on outcome (Price, Finniss et al., 2008), and a similar situation is  
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26 295 likely in the context of a clinical trial. However it seems problematic to reliability identify the  
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28 296 relevant features of the experimenter and how they would be perceived by individual  
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30 297 participants. Additionally, it seems undesirable that a hearing scientist should strive to  
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32 298 minimise warmth, empathy or friendliness towards participants, or take steps to minimise  
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34 299 his or her prestige in the eyes of the participant. Besides lowering response rates and  
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36 300 increasing dropout rates, this strategy may have the unintended consequence of having an  
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38 301 adverse impact on the trial via negative expectation (i.e. by invoking a nocebo effect).  
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44 302 It may be helpful to formulate instructions and procedures so as to minimise the effect of  
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46 303 expectation. However, this is not a perfect solution. In the present study, the only  
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48 304 information provided to participants was that the aim of the study was to evaluate hearing  
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50 305 aids with 'new technology', and no negative comments were made about the conventional  
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52 306 hearing aid. It seems that even this minimal information was sufficient to set up an  
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54 307 expectation in participants that the 'new' hearing aid must be better, and to influence the  
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3 308 performance of outcome measures. In actual hearing aid trials, participants are ethically  
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5 309 required to be given information about the test hearing aid and the hypotheses for the trial.  
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8 310 This may have a powerful effect on participant expectation (Benedetti, 2009), so that  
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10 311 placebo effects in real hearing aid trials may be larger than those reported here.  
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13 312 Dawes et al (2011) suggested that a potential method of mitigating placebo effects in  
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15 313 hearing aid trials may be to estimate the size of placebo effect, with additional effects likely  
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17 314 to be due to the experimental manipulation in question (Glasziou et al., 2007). Dawes et al  
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19 315 cautioned that this approach is problematic for two reasons: First, because placebo effects  
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21 316 are known to fluctuate in size depending on various situational variables (such as  
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23 317 experimenter characteristics) they are difficult predict. Second, the placebo effects  
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25 318 observed in this study and in Dawes et al (2011) are similar in size to experimental effects  
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27 319 typically sought in clinical hearing aid trials. It is unlikely that the experimental effect in a  
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29 320 clinical hearing aid trial would be substantially larger than that which may be due to a  
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31 321 placebo effect. Compared with unaided listening, hearing aids provide a substantial benefit.  
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33 322 Innovations in hearing aid technology are not expected to substantially increase  
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35 323 performance, and so improvements in speech recognition of a few per cent are realistic  
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37 324 goals for hearing aid researchers. For example, Wood and Lutman's (2004) hearing aid trial  
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39 325 reported statistically significantly better speech recognition of between 1 to 4 % for an  
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41 326 analogue compared to a digital hearing aid, while Valente et al's (1998) study reported a  
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43 327 1.6% advantage for the test hearing aid. In comparison, the size of the placebo effect in the  
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45 328 present study and in Dawes et al (2011) was 2-4%. Neither Wood and Lutman's nor Valente  
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47 329 et al's study used double-blinded methodology, so the apparent benefits reported in each  
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49 330 study could potentially be wholly accounted for by a placebo effect.  
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3 331 A further method of controlling for placebo effects centres on identifying and removing  
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5 332 individual placebo responders either before commencing a trial or during a 'run-in' phase. In  
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7 333 a placebo run-in, all eligible participants are given a placebo treatment. Those who respond  
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9 334 to the placebo treatment are then withdrawn from the study prior to random allocation to  
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11 335 treatment condition (For example, Gong et al, 1996). Alternatively, investigators may  
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13 336 attempt to exclude participants based on psychological predictors of the placebo response,  
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15 337 such as motivation (Geers et al., 2005), suggestibility (De Pascalis et al., 2002) or social  
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17 338 acquiescence (McNair et al., 1979). Other participant characteristics might also be used to  
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19 339 exclude placebo responders, such as those with a particular profile of symptoms or those  
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21 340 who have not previously had any experience with a particular type of treatment (Newcorn  
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23 341 et al., 2009). These strategies depend on the assumptions that there is a group of  
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25 342 participants who are reliable placebo responders across trials, and that these placebo  
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27 343 responders may be identified by certain psychological traits or demographic characteristics.  
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29 344 Neither of these assumptions seems to be supported. First, participants vary in the extent to  
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31 345 which they show a placebo response depending on a wide range of contextual and personal  
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33 346 factors (Price, Finniss et al., 2008) such that a participant may not exhibit a placebo  
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35 347 response in one situation but exhibit a placebo response in another similar situation  
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37 348 (Beecher, 1955; Liberman, 1964). Thompson (2000) concluded that because of this  
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39 349 inconsistency, there is no evidence for the effectiveness of the 'placebo run-in' as a method  
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41 350 for controlling for placebo effects. Second, the defining traits of placebo responders vary  
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43 351 between studies, and reviewers have concluded that participant characteristics are not  
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45 352 good predictors of placebo responses (Beecher, 1955; Liberman, 1964; Shapiro & Shapiro,  
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47 353 1984; Thompson, 2000; Price, Finniss et al., 2008). In relation to hearing aid trials, there is  
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49 354 no evidence to suggest whether a strategy of identifying and excluding placebo responders  
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3 355 would be feasible or not. There are several questions that must be answered. It is currently  
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5 356 unknown whether placebo effects in hearing aid trials are reliable at an individual level and  
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7 357 if there are any features that reliably characterise placebo responders that would allow  
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10 358 identification and exclusion. Even if it were possible to identify and exclude placebo  
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12 359 responders from hearing aid trials, this seems potentially problematic if the consequence  
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14 360 was the introduction of a selection bias. Excluding a specific portion of participants may  
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17 361 reduce the generalizability of the result.

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20 362 Finally, in clinical medicine placebo effects could be (and are) ethically utilised to optimise  
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22 363 the outcome for patients (Turner et al., 1994; Thompson, 2000; Price, Finniss et al., 2008).  
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24 364 Placebo effects may also be utilised in clinical audiological settings to increase patient  
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27 365 benefit. In relation to tinnitus, Tyler et al (2001) suggested that ensuring that one is  
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29 366 perceived by patients as a sympathetic, competent and confident professional may translate  
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31 367 into positive patient expectations and a higher likelihood of a successful outcome from  
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33 368 treatment. Boosting outcomes via conscious encouragement of placebo effects via  
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36 369 supportive and positive clinical interactions is likely to be ethically acceptable. Whether in  
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38 370 some circumstances and with appropriate limitations, it may be ethically acceptable to  
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40 371 prescribe a treatment primarily to elicit a placebo effect is a more difficult question (Finniss,  
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42 372 Kaptchuk et al., 2011). The utility and ethical application of placebo effects in clinical  
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44 373 audiological settings is a matter for further enquiry and informed discussion among  
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47 374 audiologists and professional bodies.

## 51 375 **Conclusions**

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55 376 Placebo effects appear to have a reliable impact on measures of hearing aid benefit such as  
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57 377 those typically used in hearing aid trials. Placebo effects need to be controlled for in hearing  
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3 378 aid trials, and given the current state of knowledge there is no satisfactory alternative to a  
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5 379 double-blind design. Hearing aid trials that do not include such controls should be  
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8 380 interpreted with caution. The influence of expectation is likely to impact on outcomes in  
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10 381 clinical audiology practice but this awaits investigation.

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Dawes *Reliability of placebo effects*

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Table 1. Coupler gain targets and measured coupler gain (in decibels) for the test hearing aids

	Frequency (Hz)					
	500	750	1000	2000	3000	4000
NAL-NL1 gain targets at 65 dB SPL	6	12	16	28	32	30
Hearing aid 1 mean gain (SD)*	6 (0.0)	9 (1.0)	14 (0.0)	23 (0.6)	32 (0.0)	34 (0.6)
Hearing aid 2 mean gain (SD)*	6 (0.0)	9 (0.6)	14 (0.0)	23 (0.6)	32 (0.0)	34 (0.6)

\* Mean gain and standard deviation is the product of three repeated coupler measures

Table 2. Four Alternative Auditory Feature (FAAF) test results

	Mean FAAF % correct 'New' (SD)	Mean FAAF % correct 'Conventional' (SD)	Difference (New – Conventional)	Z*	p	Effect size (r)
Current study	70.9 (12.6)	66.8 (14.1)	4.1%  (95% CI: 0.6 to 7.6)	2.30	0.02	0.15
Dawes et al 2011	62.3 (10.4)	60.7 (9.0)	1.6% (95% CI: -1.0 to 4.2)	1.84	0.06	0.08

\*Wilcoxon signed-rank test statistic.



Table 3. Sound Quality Ratings

		Comfort	Clarity	Overall Impression	Overall sound quality rating
Current study	Mean rating "New" (SD)	7.98 (1.76)	8.77 (0.93)	7.84 (1.86)	8.20 (1.39)
	Mean rating "Conventional" (SD)	7.22 (1.62)	7.76 (1.27)	7.16 (1.62)	7.38 (1.37)
	Z*	-2.9	-3.14	-2.36	-2.99
	p (two tailed)	0.001	0.004	0.018	0.003
	Effect size (r)	0.22	0.41	0.19	0.28
Dawes et al 2011	Mean rating 'new' (SD)	8.95 (1.12)	9.28 (1.15)	9.00 (1.11)	9.12 (1.02)
	Mean rating 'conventional' (SD)	8.40 (1.21)	8.61 (1.28)	8.1 (1.47)	8.35 (1.17)
	Z*	-1.94	-2.77	-2.98	-2.88
	p (two tailed)	0.053	0.006	0.003	0.004
	Effect size (r)	0.23	0.27	0.33	0.33

\*Wilcoxon signed-rank test statistic.