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CONSTRAINED DYNAMIC OPTIMALITY AND BINOMIAL TERMINAL WEALTH*

J. L. PEDERSEN[†] AND G. PESKIR[‡]

Abstract. We assume that the wealth process X^u is self-financing and generated from the initial wealth by holding a fraction u of X^u in a risky stock (whose price follows a geometric Brownian motion) and the remaining fraction $1-u$ of X^u in a riskless bond (whose price compounds exponentially with interest rate $r \in \mathbb{R}$). Letting $\mathbb{P}_{t,x}$ denote a probability measure under which X^u takes value x at time t , we study the dynamic version of the nonlinear optimal control problem $\inf_u \text{Var}_{t, X_t^u}(X_T^u)$ where the infimum is taken over admissible controls u subject to $X_t^u \geq e^{-r(T-t)}g$ and $\mathbb{E}_{t, X_t^u}(X_T^u) \geq \beta$ for $t \in [0, T]$. The two constants g and β are assumed to be given exogenously and fixed. By conditioning on the expected terminal wealth value, we show that the nonlinear problem can be reduced to a family of linear problems. Solving the latter using a martingale method combined with Lagrange multipliers, we derive the dynamically optimal control u_*^d in closed form and prove that the dynamically optimal terminal wealth X_T^d can only take two values g and β . This binomial nature of the dynamically optimal strategy stands in sharp contrast with other known portfolio selection strategies encountered in the literature. A direct comparison shows that the dynamically optimal (time-consistent) strategy outperforms the statically optimal (time-inconsistent) strategy in the problem.

Key words. constrained nonlinear optimal control, dynamic optimality, static optimality, mean-variance analysis, martingale, Lagrange multiplier, geometric Brownian motion, Markov process

AMS subject classifications. Primary, 60H30, 60J65; Secondary, 91G80, 93E20

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1. Introduction. Imagine an investor who has an initial wealth which he wishes to exchange between a *risky* stock and a *riskless* bond, in a self-financing manner, dynamically in time, so as to *minimize his risk* in obtaining a desired return at the given terminal time. In line with the mean-variance analysis of Markowitz [14], where the optimal portfolio selection problem of this kind was solved in a single period model (see, e.g., Merton [15] and the references therein), we will identify the return with the expectation of the terminal wealth, and the risk with the variance of the terminal wealth. The quadratic nonlinearity of the variance then moves the resulting optimal control problem outside the scope of the standard optimal control theory, which yields optimal controls dependent on the initial time/wealth, and hence *time inconsistent* (see [16] for further details and a historical review dating back to Strotz [20]).

A new methodology for solving nonlinear control problems of this kind has been recently developed in [16]. Its central idea rests on the concept of *dynamic optimality* which consists of continuous rebalancing of optimal controls upon overruling all the past controls. This makes *dynamically optimal* controls *time consistent*. It stands in contrast with the *statically optimal* (precommitted) controls addressed above which are *time inconsistent*. For more details, including comparisons with the subgame-perfect Nash equilibrium controls (dating back to the concept of “consistent planning” in Strotz [20]) we refer the reader to [16] and the references therein.

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Assuming that the stock price follows a geometric Brownian motion and the bond price compounds exponentially, in [16] we derived the dynamically optimal controls for the investor aiming to minimize the variance of his/her terminal wealth X_T^u over all admissible controls u (representing the fraction of the wealth held in the stock) such that the expectation of X_T^u is bounded below by a given constant β . We showed in [16, Corollary 7] that the dynamically optimal wealth process X^d solves a meander-type equation which makes X^d hit β exactly at the terminal time T (not before). This was established without any pathwise constraints on the wealth process, which could take low/negative values of unlimited size, leading to a potential bankruptcy, and thus making the trading strategy less attractive to the investor.

Motivated by this deficiency, in this paper we consider the analogous variance minimizing problem obtained upon imposing the guarantee (stop-loss rule) that the wealth process always stays above a given constant g , regardless of whether the investment is favorable or not. In addition to finance, where $g = 0$ corresponds to bankruptcy, such constraints are also widely used in economics (see, e.g., [3] and the references therein). By conditioning on the expected terminal wealth value we show that the nonlinear problem can be reduced to a family of linear problems. Solving the latter using a martingale method (cf. Pliska [17]) combined with Lagrange multipliers (cf. White [21]) we derive the dynamically optimal control u_*^d in closed form and establish that the dynamically optimal wealth process X^d can be characterized as the unique (strong) solution to a stochastic differential equation with time-dependent coefficients. By analyzing this stochastic differential equation (extending Feller's test to time-inhomogeneous diffusions that is of independent interest) we prove that the dynamically optimal terminal wealth X_T^d can only take two values g and β . This *binomial* nature of the dynamically optimal strategy stands in sharp contrast with other known portfolio selection strategies encountered in the literature. A direct comparison shows that the dynamically optimal (time-consistent) strategy outperforms the statically optimal (time-inconsistent) strategy in the minimizing variance problem.

The portfolio selection problem of minimizing the variance of the terminal wealth given its expected value was first studied by Richardson [18] without imposing pathwise constraints on low/negative wealth. Lagrange reformulations of this problem were studied in [12] (discrete time) and [22] (continuous time) also without imposing pathwise constraints on low/negative wealth (see also [13]). Korn and Trautmann [11] studied Richardson's problem under a pathwise constraint on low/negative wealth using duality techniques (see also [7] and [10]). Bielecki et al. [5] extended these results to a more general setting using backward stochastic differential equations (see also [4] for applications in insurance mathematics).

The portfolio selection strategies in all these papers are, as pointed out by Basak and Chabakauri [1], precommitted (statically optimal) and hence time inconsistent (see also [2]). The dynamically optimal strategies derived in the present paper are, to our knowledge, the first known optimal strategies that are time consistent in the variance minimizing problem with pathwise constraints on low/negative wealth. It appears to be a challenging problem to derive such strategies in the context of the subgame-perfect Nash equilibrium, and any comparison with other time-consistent strategies is currently unavailable.

2. Formulation of the problem. Assume that the *riskless* bond price B solves

$$(2.1) \quad dB_t = r B_t dt,$$

with $B_0 = b$ for some $b > 0$ where $r \in \mathbb{R}$ is the interest rate. Let the *risky* stock price S follow a geometric Brownian motion solving

$$(2.2) \quad dS_t = \mu S_t dt + \sigma S_t dW_t$$

with $S_0 = s$ for some $s > 0$, where $\mu \in \mathbb{R}$ is the drift, $\sigma > 0$ is the volatility, and W is a standard Brownian motion defined on a probability space $(\Omega, \mathcal{F}, \mathbb{P})$. Note that a unique solution to (2.1) is given by $B_t = be^{rt}$ and recall that a unique strong solution to (2.2) is given by $S_t = s \exp(\sigma W_t + (\mu - (\sigma^2/2))t)$ for $t \geq 0$.

Consider the investor who has an initial wealth $x_0 \in \mathbb{R}$ which he/she wishes to exchange between B and S in a *self-financing* manner (with no exogenous infusion or withdrawal of wealth) dynamically in time up to the given horizon $T > 0$. It is then well known (see, e.g., [6, Chapter 6]) that the investor's wealth X^u solves

$$(2.3) \quad dX_t^u = (r(1-u_t) + \mu u_t) X_t^u dt + \sigma u_t X_t^u dW_t,$$

with $X_{t_0}^u = x_0$ where u_t denotes the fraction of the investor's wealth held in the stock at time $t \in [t_0, T]$ for $t_0 \in [0, T]$ given and fixed. Note that (i) $u_t < 0$ corresponds to *short selling* of the stock, (ii) $u_t > 1$ corresponds to *borrowing* from the bond, and (iii) $u_t \in [0, 1]$ corresponds to a *long position* in both the stock and the bond.

To simplify the exposition, we will assume that the control u in (2.3) is given by $u_t = u(t, X_t^u)$, where $(t, x) \mapsto u(t, x) \cdot x$ is a continuous function from $[0, T] \times \mathbb{R}$ into \mathbb{R} for which the stochastic differential equation (2.3) understood in Itô's sense, has a unique *strong* solution X^u (meaning that the solution X^u to (2.3) is adapted to the natural filtration of W , and if \tilde{X}^u is another solution to (2.3) of this kind, then X^u and \tilde{X}^u are equal almost surely). In what follows, we will call controls of this kind *admissible*. For more general controls, yielding the same solution to the problem to be presented below, we refer the reader to [16, section 2] and the references therein. Recalling that the natural filtration of S coincides with the natural filtration of W , we see that admissible controls have a natural financial interpretation, as they are obtained as deterministic (measurable) functionals of the observed stock price. We will denote the natural filtration by $(\mathcal{F}_t)_{0 \leq t \leq T}$ in what follows (omitting superscripts S and W for simplicity).

For a given admissible control u we let $\mathbb{P}_{t,x}$ denote the probability measure under which the solution X^u to (2.3) takes value x at time t for $(t, x) \in [0, T] \times \mathbb{R}$. Note that X^u is a (strong) Markov process with respect to $\mathbb{P}_{t,x}$ for $(t, x) \in [0, T] \times \mathbb{R}$. To formulate the investor's problem, we consider admissible controls u obeying the pathwise constraint

$$(2.4) \quad X_s^u \geq e^{-r(T-s)} g \quad \mathbb{P}_{t,x}\text{-a.s. for all } s \in [t, T],$$

where g is a given constant playing the financial role of a *guarantee* (not to be strictly crossed from above). This is a *primary constraint* on the investor's wealth process. Note that if $x = e^{-r(T-t)} g$, then the only admissible control u which can meet the primary constraint (2.4) equals 0 (note also that if $x < e^{-r(T-t)} g$, then no admissible control can meet the constraint). This corresponds to investing all the wealth in the bond over the remaining time interval $[t, T]$. If $x \in (e^{-r(T-t)} g, e^{-r(T-t)} \beta)$ for a given constant β (strictly larger than g), we also impose a *secondary constraint* on the investor's wealth process, which is expressed as

$$(2.5) \quad \mathbb{E}_{t,x}(X_T^u) \geq \beta,$$

where β plays the financial role of a *target* (to be met at the terminal time T). Note that if $x \geq e^{-r(T-t)}\beta$, then the admissible control u which meets the primary constraint (2.4) and the secondary constraint (2.5) equals 0. This again corresponds to investing all the wealth in the bond with zero variance (risk) at the terminal time T . For these reasons we will see from the problem formulation below that the only interesting/nontrivial case is when $x \in (e^{-r(T-t)}g, e^{-r(T-t)}\beta)$ for $(t, x) \in [0, T] \times \mathbb{R}$. Given that we are interested in admissible controls yielding minimal terminal risk (identified with variance), there is no loss of generality in assuming that the terminal wealth of each admissible control is square-integrable. A sufficient condition for the latter expectation to be finite is that $\mathbb{E}_{t,x}[\int_t^T (1+u_s^2)(X_s^u)^2 ds] < \infty$ for $(t, x) \in [0, T] \times \mathbb{R}$. The imposed requirements may now be summarized as follows.

DEFINITION 1 (pathwise constrained strategies). *The family of admissible strategies u satisfying the primary constraint (2.4) with the secondary constraint (2.5) (whenever feasible) and the integrability condition $\mathbb{E}_{t,x}[\int_t^T (1+u_s^2)(X_s^u)^2 ds] < \infty$ for $(t, x) \in [0, T] \times \mathbb{R}$ will be denoted by $C_{g,\beta}$, where $g < \beta$ are given and fixed constants.*

Having defined pathwise constrained strategies, consider the optimal control problem

$$(2.6) \quad V(t, x) = \inf_{u \in C_{g,\beta}} \text{Var}_{t,x}(X_T^u)$$

for $(t, x) \in [0, T] \times \mathbb{R}$, with $g < \beta$ given and fixed. Recall that we can focus on the case when $x \in (e^{-r(T-t)}g, e^{-r(T-t)}\beta)$ for $(t, x) \in [0, T] \times \mathbb{R}$, as other cases are evident. The problem (2.6) seeks to minimize the investor’s risk (identified with the variance of X_T^u), upon applying a control u which yields a desired investor’s return (identified with the expectation of X_T^u), and keeps the investor’s wealth above a desired/guaranteed threshold at all times by the terminal time T . This identification is done in line with the mean-variance analysis of Markowitz [14]. A (time-consistent) solution to the problem (2.6) is known in the pathwise unconstrained case corresponding to no guarantee (see [16, Corollary 7]). The main purpose of the present paper is to derive a (time-consistent) solution in the pathwise constrained case.

Due to the quadratic nonlinearity of the second term in the expression $\text{Var}_{t,x}(X_T^u) = \mathbb{E}_{t,x}[(X_T^u)^2] - [\mathbb{E}_{t,x}(X_T^u)]^2$ it is known that the problem (2.6) falls outside the scope of the standard/linear optimal control theory for Markov processes. In addition to the static (precommitted) formulation of the nonlinear problem (2.6), where the minimization takes place relative to the initial point (t, x) which is given and fixed (cf. [16, Definition 1]), one is also naturally led to consider a dynamic formulation of the nonlinear problem (2.6), in which each new position of the controlled process $((t, X_t^u))_{t \in [0, T]}$ yields a new optimal control problem, to be solved upon overruling all the past problems (cf. [16, Definition 2]).

For the sake of completeness we now recall the definition of dynamic optimality for pathwise constrained strategies in $C_{g,\beta}$. Recall that all controls throughout refer to admissible controls as defined/discussed above.

DEFINITION 2 (dynamic optimality). *A control u_* is dynamically optimal in (2.6), if for every given and fixed $(t, x) \in [0, T] \times \mathbb{R}$ and every control v from $C_{g,\beta}$ such that $v(t, x) \neq u_*(t, x)$, there exists a control w in $C_{g,\beta}$ satisfying $w(t, x) = u_*(t, x)$ such that*

$$(2.7) \quad \text{Var}_{t,x}(X_T^w) < \text{Var}_{t,x}(X_T^v).$$

Dynamic optimality above is understood in the “strong” sense. Replacing the strict inequality in (2.7) by an inequality would yield dynamic optimality in the “weak” sense.

Note that the dynamic optimality corresponds to solving infinitely many optimal control problems dynamically in time, where each new position of the controlled process $((t, X_t^u))_{t \in [0, T]}$ yields a new optimal control problem, to be solved upon overruling all the past problems. In effect, this corresponds to continuous rebalancing of optimal controls, upon overruling all the past controls, which makes the dynamically optimal controls time consistent (cf. [16]).

3. Solution to the problem. To formulate the main results below we introduce the following notation:

$$(3.1) \quad L_t = e^{-\delta W_t - \frac{\delta^2}{2}t} \quad \text{and} \quad H_t = e^{-rt} L_t,$$

$$(3.2) \quad F_0(T-t, \lambda) := E[(\lambda - H_{T-t})^+] = \lambda \varphi\left(\frac{\log \lambda + (r + \frac{\delta^2}{2})(T-t)}{|\delta|\sqrt{T-t}}\right) \\ - e^{-r(T-t)} \Phi\left(\frac{\log \lambda + (r - \frac{\delta^2}{2})(T-t)}{|\delta|\sqrt{T-t}}\right),$$

$$(3.3) \quad F_1(T-t, \lambda) = E[H_{T-t}(\lambda - H_{T-t})^+] = \lambda e^{-r(T-t)} \varphi\left(\frac{\log \lambda + (r - \frac{\delta^2}{2})(T-t)}{|\delta|\sqrt{T-t}}\right) \\ - e^{(\delta^2 - 2r)(T-t)} \Phi\left(\frac{\log \lambda + (r - \frac{3\delta^2}{2})(T-t)}{|\delta|\sqrt{T-t}}\right),$$

$$(3.4) \quad F_2(T-t, \lambda) := E[H_{T-t}^2 I(\lambda > H_{T-t})] \\ = e^{(\delta^2 - 2r)(T-t)} \Phi\left(\frac{\log \lambda + (r - \frac{3\delta^2}{2})(T-t)}{|\delta|\sqrt{T-t}}\right),$$

$$(3.5) \quad F(T-t, \lambda) := \frac{F_1(T-t, \lambda)}{F_0(T-t, \lambda)}$$

for $t \in [0, T]$ and $\lambda > 0$, where the expressions on the right-hand sides of (3.2)–(3.4) are derived using standard arguments (recall that $\Phi(x) = (1/\sqrt{2\pi}) \int_{-\infty}^x e^{-y^2/2} dy$ denotes the standard normal distribution function for $x \in \mathbb{R}$).

THEOREM 3. Consider the optimal control problem (2.6), where X^u solves (2.3) with $X_{t_0}^u = x_0$ under \mathbb{P}_{t_0, x_0} for $(t_0, x_0) \in [0, T] \times \mathbb{R}$ given and fixed. Recall that B solves (2.1), S solves (2.2), and we set $\delta = (\mu - r)/\sigma$ for $\mu \in \mathbb{R}$, $r \in \mathbb{R}$, and $\sigma > 0$. Throughout we assume that $\delta \neq 0$ to exclude the trivial case (the case $\delta = 0$ follows by passage to the limit when the nonzero δ tends to 0).

(A) The dynamically optimal control u_*^d in (2.6) is given by

$$(3.6) \quad u_*^d(t, x) = \frac{\delta}{\sigma} \frac{1}{x} (x - e^{-r(T-t)} g) a(T-t, x),$$

where the function a is defined by

$$(3.7) \quad a(T-t, x) = \frac{F_2(T-t, \lambda(T-t, x))}{F_1(T-t, \lambda(T-t, x))}$$

and $\lambda = \lambda(T-t, x)$ is the unique strictly positive solution to the equation

$$(3.8) \quad F(T-t, \lambda) = \frac{x - e^{-r(T-t)}g}{\beta - g}$$

for $(t, x) \in [t_0, T] \times \mathbb{R}$ such that $e^{-r(T-t)}g < x < e^{-r(T-t)}\beta$. If $x = e^{-r(T-t)}g$ or $x \geq e^{-r(T-t)}\beta$, then $u_*^d(t, x) = 0$ (as discussed prior to Definition 1 above).

(B) The dynamically optimal controlled process X^d in (2.6) can be characterized as the unique strong solution to the stochastic differential equation

$$(3.9) \quad \begin{aligned} dX_t^d &= [rX_t^d + \delta^2(X_t^d - e^{-r(T-t)}g) a(T-t, X_t^d)] dt \\ &\quad + \delta(X_t^d - e^{-r(T-t)}g) a(T-t, X_t^d) dW_t \end{aligned}$$

satisfying $X_{t_0}^d = x_0 \in (e^{-r(T-t_0)}g, e^{-r(T-t_0)}\beta)$, until the first exit time of X_t^d from the interval $(e^{-r(T-t)}g, e^{-r(T-t)}\beta)$, when t runs from t_0 to T , followed by the constant motion (at its left-hand boundary) corresponding to $u_*^d = 0$ afterwards. Setting

$$(3.10) \quad \tau_\beta = \inf \{ t \in [t_0, T] \mid X_t^d = e^{-r(T-t)}\beta \},$$

$$(3.11) \quad \tau_g = \inf \{ t \in [t_0, T] \mid X_t^d = e^{-r(T-t)}g \},$$

the following (binomial) relations are satisfied:

$$(3.12) \quad \mathbf{P}_{t_0, x_0}(\tau_\beta < T) = 0 \quad \text{and} \quad \mathbf{P}_{t_0, x_0}(\tau_\beta = T) > 0,$$

$$(3.13) \quad \mathbf{P}_{t_0, x_0}(\tau_g < T) > 0.$$

This means that the dynamically optimally controlled process X^d (i) hits the target β exactly at the terminal time T (never before T) with a strictly positive probability, and (ii) hits the guarantee g strictly before the terminal time T with a strictly positive probability (see Figure 1).

Proof. We assume throughout that the process X^u solves the stochastic differential equation (2.3), with $X_{t_0}^u = x_0$ under \mathbf{P}_{t_0, x_0} for $(t_0, x_0) \in [0, T] \times \mathbb{R}$ given and fixed, where u is any admissible control as defined/discussed above.

(A) Recall that the optimal control problem (2.6) seeks to minimize

$$(3.14) \quad \text{Var}_{t_0, x_0}(X_T^u) = \mathbf{E}_{t_0, x_0}[(X_T^u)^2] - [\mathbf{E}_{t_0, x_0}(X_T^u)]^2$$

over all admissible controls u subject to the primary (pathwise) constraint

$$(3.15) \quad X_t^u \geq e^{-r(T-t)}g \quad \mathbf{P}_{t_0, x_0}\text{-a.s. for all } t \in [t_0, T]$$

and the secondary (mean value) constraint

$$(3.16) \quad \mathbf{E}_{t_0, x_0}(X_T^u) \geq \beta$$

as specified in Definition 1 above.

1. Using (3.1) and passing to the equivalent martingale measure defined by

$$(3.17) \quad d\tilde{\mathbf{P}}_{t_0, x_0} := (L_T/L_{t_0}) d\mathbf{P}_{t_0, x_0},$$

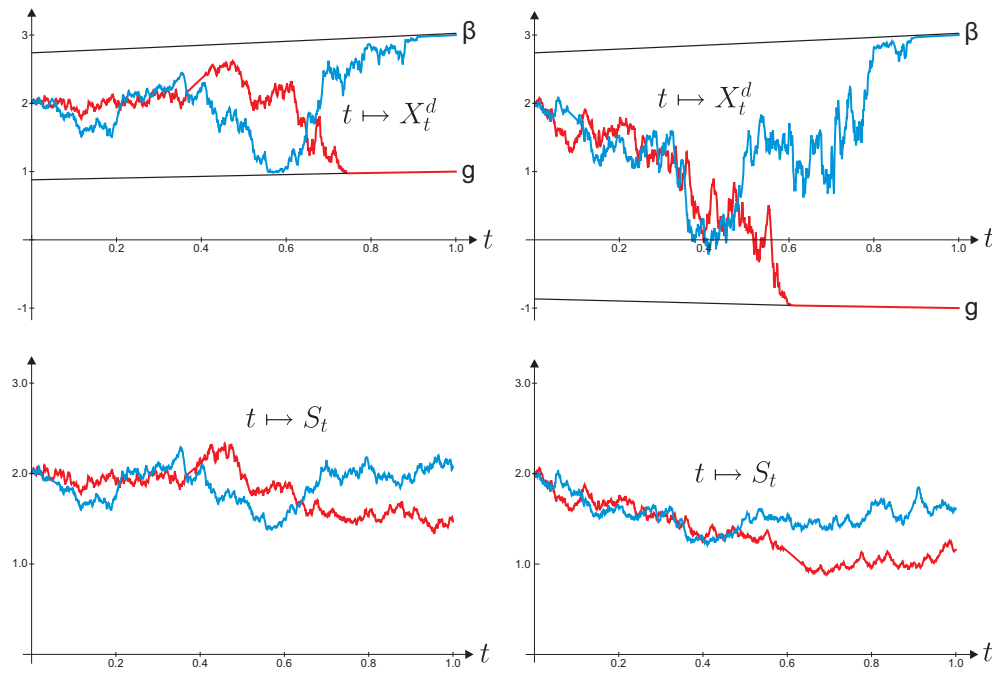


FIG. 1. The dynamically optimal wealth $t \mapsto X_t^d$ in the constrained problem (2.6) of Theorem 3 obtained from the stock price $t \mapsto S_t$ when $t_0 = 0$, $x_0 = 2$, $S_0 = 2$, $\beta = 3$, $r = 0.1$, $\mu = 0$, $\sigma = 0.4$, $T = 1$, $g = 1$ (left-hand graphs), and $g = -1$ (right-hand graphs). Note that the expected value of S_T equals 2, which is strictly smaller than β .

under which $\tilde{W}_t := W_t - W_{t_0} + \delta(t - t_0)$ is a standard Brownian motion for $t \in [t_0, T]$, and recalling that the discounted wealth process defined by

$$(3.18) \quad \tilde{X}_t^u := e^{-r(t-t_0)} X_t^u$$

solves the stochastic differential equation

$$(3.19) \quad d\tilde{X}_t^u = \sigma u_t \tilde{X}_t^u d\tilde{W}_t$$

and defines a martingale for $t \in [t_0, T]$, one sees that the pathwise constraint (3.15) is equivalent to the (seemingly weaker) terminal constraint

$$(3.20) \quad X_T^u \geq g \quad \mathbb{P}_{t_0, x_0}\text{-a.s.}$$

Indeed, while (3.15) evidently implies (3.20), to see that the converse is also true, it is sufficient to apply the martingale property of \tilde{X}^u (cf. [16, Remark 9]) which by means of (3.20) yields

$$(3.21) \quad e^{-r(t-t_0)} X_t^u = \tilde{X}_t^u = \tilde{\mathbb{E}}_{t_0, x_0}(\tilde{X}_T^u | \mathcal{F}_t) \geq e^{-r(T-t_0)} g$$

and implies (3.15) as claimed. By the martingale property of \tilde{X}^u we also have

$$(3.22) \quad \tilde{\mathbb{E}}_{t_0, x_0}(\tilde{X}_T^u) = x_0,$$

which can be rewritten using (3.17) and (3.18) as

$$(3.23) \quad \mathbb{E}_{t_0, x_0}[(H_T/H_{t_0})X_T^u] = x_0,$$

showing that admissible controls u stand in one-to-one correspondence with \mathcal{F}_T -measurable square-integrable random variables X_T^u satisfying (3.23). In the latter conclusion we use the Itô–Clark theorem (see, e.g., [9, pp. 185–189]) applied to the stochastic integral with respect to \tilde{W} over $[t_0, T]$ under $\tilde{\mathbb{P}}_{t_0, x_0}$. Details of this will be utilized further in the proof below.

2. The previous considerations show that the optimal control problem (3.14) is equivalent to the pointwise constrained optimization problem

$$(3.24) \quad \inf_{X \geq g} \mathbb{E}(X^2)$$

subject to the mean value constraints

$$(3.25) \quad \mathbb{E}(X) = \beta \quad \text{and} \quad \mathbb{E}(HX) = x,$$

where we impose equality in (3.16) and simplify the notation by setting $X := X_T^u$, $\mathbb{P} := \mathbb{P}_{t_0, x_0}$, $H := H_T/H_{t_0}$, and $x = x_0$ (we will return to the original notation below).

3. To solve the problem (3.24)–(3.25) we optimize the Lagrangian

$$(3.26) \quad \inf_{X \geq g} \mathbb{E}(X^2 - \lambda_1 X - \lambda_2 HX)$$

for λ_1 and λ_2 in \mathbb{R} to be determined. Pointwise minimization of the expression under the expectation sign in (3.26) then yields the (unique) solution

$$(3.27) \quad X_* = \frac{1}{2}(\lambda_1 + \lambda_2 H) \vee g = \left(\frac{\lambda_1}{2} + \frac{\lambda_2}{2} H - g\right)^+ + g = \kappa(\lambda - H)^+ + g$$

upon setting $\kappa = -\lambda_2/2$ and $\lambda = (\lambda_1/2 - g)/(-\lambda_2/2)$, where we impose $\kappa > 0$, i.e., $\lambda_2 < 0$ (see Remark 4 below for fuller details). Using (3.25) we get

$$(3.28) \quad \beta = \mathbb{E}(X_*) = \kappa \mathbb{E}[(\lambda - H)^+] + g,$$

$$(3.29) \quad x = \mathbb{E}(HX_*) = \kappa \mathbb{E}[H(\lambda - H)^+] + g \mathbb{E}(H).$$

From (3.28) and (3.29) we find that

$$(3.30) \quad \kappa = \frac{\beta - g}{\mathbb{E}[(\lambda - H)^+]} = \frac{x - g \mathbb{E}(H)}{\mathbb{E}[H(\lambda - H)^+]}$$

The second identity in (3.30) can be rewritten as

$$(3.31) \quad \frac{\mathbb{E}[H(\lambda - H)^+]}{\mathbb{E}[(\lambda - H)^+]} = \frac{x - g \mathbb{E}(H)}{\beta - g}.$$

4. We now show that there exists a unique $\lambda > 0$ solving (3.31) when $x \in (g \mathbb{E}(H), \beta \mathbb{E}(H))$ and $\mathbb{P}(0 < |H| \leq \varepsilon) > 0$ for each $\varepsilon > 0$ (assuming also that $\mathbb{E}(H)$ exists and is finite).

For this, define the function

$$(3.32) \quad F(\lambda) := \frac{\mathbb{E}[H(\lambda - H)^+]}{\mathbb{E}[(\lambda - H)^+]}$$

for $\lambda > 0$. We then claim that

$$(3.33) \quad \lambda \mapsto F(\lambda) \text{ is strictly increasing on } (0, \infty),$$

$$(3.34) \quad F(0+) = 0 \quad \text{and} \quad F(+\infty) = E(H).$$

To verify (3.33) it is enough to show that $F'(\lambda) > 0$ for all $\lambda > 0$ such that $P(H = \lambda) = 0$ (recall that there could be at most countably many $\lambda > 0$ for which this fails). Taking any such $\lambda > 0$ and differentiating under the expectation sign in (3.32) (which is easily justified using standard means) we see that $F'(\lambda) > 0$ if and only if

$$(3.35) \quad E[HI(\lambda > H)] E[(\lambda - H)I(\lambda > H)] > E[H(\lambda - H)I(\lambda > H)] P(\lambda > H),$$

which in turn is equivalent to

$$(3.36) \quad (E[HI(\lambda > H)])^2 < E[H^2I(\lambda > H)] P(\lambda > H).$$

Writing the integrand on the left-hand side as $HI(\lambda > H) = HI(\lambda > H) \cdot I(\lambda > H)$, we recognize (3.36) as the Hölder inequality. It is well known (see [8, Exc. 3.35]) that equality holds in the Hölder inequality if and only if $|H|I(\lambda > H) = cI(\lambda > H)$ for some $c \geq 0$. Hence if the strict inequality in (3.36) would not hold, this would violate the hypotheses that $P(0 < |H| \leq \varepsilon) > 0$ for all $\varepsilon > 0$. It follows therefore that (3.33) holds as claimed.

The first equality in (3.34) follows from the fact that $F(\lambda) \leq \lambda$ for all $\lambda > 0$, and the second equality in (3.34) follows by noting that $F(\lambda) = E[H(1-H/\lambda)^+]/E[(1-H/\lambda)^+]$ and using the dominated converge theorem upon letting $\lambda \rightarrow \infty$. Noting that $F(\lambda)$ equals the left-hand side in (3.31) and that $\lambda \mapsto F(\lambda)$ is continuous on $(0, \infty)$, we see by (3.33)+(3.34) that there exists a unique $\lambda > 0$ solving (3.31) as claimed.

5. The existence of a unique $\lambda > 0$ solving (3.31) combined with (3.30) shows that X_* defined in (3.27) is the solution to the constrained problem (3.24)+(3.25). Recalling that $X := X_T^u$, $P := P_{t_0, x_0}$, $H := H_T/H_{t_0}$, and $x = x_0$, we see that (3.31) is the same as (3.8), with $x = x_0$ in the statement of the theorem. We will return to this point later in the proof. Now we turn to finding the statically optimal control u_*^s and its wealth process X^s , which yields X_* at the terminal time T .

6. To find the statically optimal control u_*^s and its wealth process X^s satisfying $X_T^s = X_*$, we can exploit the martingale property of \tilde{X}^s defined as in (3.18) with $u = u_*^s$, which in view of (3.27) yields the identities

$$(3.37) \quad X_t^s = e^{-r(T-t)} \tilde{E}_{t_0, x_0}(X_* | \mathcal{F}_t) = e^{-r(T-t)} [\kappa \tilde{E}_{t_0, x_0}((\lambda - H_T/H_{t_0})^+ | \mathcal{F}_t) + g]$$

for $t \in [t_0, T]$, where $\lambda = \lambda(T - t_0, x_0)$ is the unique strictly positive solution to (3.31), and $\kappa = \kappa(T - t_0, x_0)$ is given by (3.30) with $P = P_{t_0, x_0}$, $H = H_T/H_{t_0}$, and $x = x_0$. Recalling (3.1) we see that $\tilde{H}_t := H_t/H_{t_0}$ solves

$$(3.38) \quad d\tilde{H}_t = (\delta^2 - r)\tilde{H}_t dt - \delta\tilde{H}_t \tilde{W}_t$$

and defines a geometric Brownian motion under \tilde{P}_{t_0, x_0} for $t \in [t_0, T]$. Writing H_T/H_{t_0} in (3.37) as $(H_t/H_{t_0})(H_T/H_t)$, and using independent increments of \tilde{W} , we see that (3.37) reads

$$(3.39) \quad X_t^s = e^{-r(T-t)} [\kappa P(T-t, \tilde{H}_t) + g],$$

where the function P is defined by

$$(3.40) \quad \begin{aligned} P(T-t, h) &= \tilde{E}_{t_0, x_0} [(\lambda - h(H_T/H_t))^+] = \mathbf{E}_{t_0, x_0} \left[\frac{L_T}{L_t} \frac{L_t}{L_{t_0}} (\lambda - h(H_T/H_t))^+ \right] \\ &= e^{r(T-t)} \mathbf{E} [H_{T-t} (\lambda - hH_{T-t})^+] \end{aligned}$$

for $t \in [t_0, T]$, and in the final equality we use that L_t/L_{t_0} is \mathcal{F}_t -measurable (with mean equal one) and both L_T/L_t and H_T/H_t are independent from \mathcal{F}_t . Note that the second identity in (3.3) can be used to express $P(T-t, h)$ in closed form using the standard normal distribution function Φ ; however, we will omit this explicit expression for simplicity.

From (3.39) we see that

$$(3.41) \quad Z_t^s := e^{r(T-t)} X_t^s = \kappa P(T-t, \tilde{H}_t) + g$$

for $t \in [t_0, T]$. Applying Itô's formula to the right-hand side of (3.41) and using that Z^s is a martingale, we find by means of (3.38) and (3.19) that

$$(3.42) \quad dZ_t^s = \kappa \frac{\partial P}{\partial h}(T-t, \tilde{H}_t) (-\delta) \tilde{H}_t d\tilde{W}_t = \sigma u_t^s Z_t^s d\tilde{W}_t$$

for $t \in [t_0, T]$. It follows from the second identity in (3.42) that

$$(3.43) \quad \begin{aligned} u_t^s &= -\frac{\delta}{\sigma} \kappa \frac{\partial P}{\partial h}(T-t, \tilde{H}_t) \frac{\tilde{H}_t}{Z_t^s} = -\frac{\delta}{\sigma} \frac{x_0 - e^{-r(T-t_0)}g}{\mathbf{E}[\tilde{H}_T(\lambda - \tilde{H}_T)^+]} \frac{\partial P}{\partial h}(T-t, \tilde{H}_t) \frac{\tilde{H}_t}{Z_t^s} \\ &= -\frac{\delta}{\sigma} \frac{e^{r(t-t_0)}x_0 - e^{-r(T-t)}g}{P(T-t_0, 1)} \frac{\partial P}{\partial h}(T-t, \tilde{H}_t) \frac{\tilde{H}_t}{X_t^s} \end{aligned}$$

for $t \in [t_0, T]$. Note that the second identity in (3.4) can be used to express $(\partial P/\partial h)(T-t, h)$ in closed form using the standard normal distribution function Φ ; however, we will omit the explicit expression for simplicity. Note also that $h \mapsto P(T-t, h)$ is strictly increasing (and continuous) so that from (3.39) we see that

$$(3.44) \quad \begin{aligned} \tilde{H}_t &= P_h^{-1}(T-t, (e^{r(T-t)}X_t^s - g)/\kappa) \\ &= P_h^{-1}\left(T-t, \frac{e^{r(T-t)}X_t^s - g}{x_0 - e^{-r(T-t_0)}g} \mathbf{E}[\tilde{H}_T(\lambda - \tilde{H}_T)^+]\right) \\ &= P_h^{-1}\left(T-t, \frac{e^{r(T-t)}X_t^s - g}{e^{r(T-t_0)}x_0 - g} P(T-t_0, 1)\right) \end{aligned}$$

for $t \in [t_0, T]$. Inserting \tilde{H}_t into the right-hand side of (3.43), we find that the statically optimal control in (2.6) is given by

$$(3.45) \quad \begin{aligned} u_*^s(t, x) &= -\frac{\delta}{\sigma} \frac{1}{x} \frac{e^{r(t-t_0)}x_0 - e^{-r(T-t)}g}{P(T-t_0, 1)} \\ &\quad \times \frac{\partial P}{\partial h}\left(T-t, P_h^{-1}\left(T-t, \frac{e^{r(T-t)}x - g}{e^{r(T-t_0)}x_0 - g} P(T-t_0, 1)\right)\right) \\ &\quad \times P_h^{-1}\left(T-t, \frac{e^{r(T-t)}x - g}{e^{r(T-t_0)}x_0 - g} P(T-t_0, 1)\right) \end{aligned}$$

for $(t, x) \in [t_0, T] \times \mathbb{R}$ such that $e^{-r(T-t)}g < x < e^{-r(T-t)}\beta$. (If $x = e^{-r(T-t)}g$ or $x \geq e^{-r(T-t)}\beta$, then $u_*^d(t, x) = 0$ as discussed prior to Definition 1 above.)

7. The statically optimal wealth process X^s can be characterized as the unique strong solution to the stochastic differential equation (2.3) obtained by inserting the statically optimal control $u_t^s = u_*^s(t, X_t^s)$ in place of u_t for $t \in [t_0, T]$. From (3.27) we see that

$$(3.46) \quad X_T^s = \kappa(\lambda - H)^+ + g,$$

where $\lambda = \lambda(T - t_0, x_0)$ is the unique strictly positive solution to (3.31) and $\kappa = \kappa(T - t_0, x_0)$ is given by (3.30) with $\mathbb{P} = \mathbb{P}_{t_0, x_0}$, $H = H_T/H_{t_0}$, and $x = x_0$. The static value function is given by

$$(3.47) \quad V_s(t_0, x_0) = \text{Var}_{t_0, x_0}(X_T^s) = (\kappa(t_0, x_0))^2 \text{Var}[(\lambda(t_0, x_0) - H_{T-t_0})^+].$$

Note that the latter variance can be readily calculated; however, we will omit the explicit expression for simplicity.

8. To obtain a candidate for the dynamically optimal control u_*^d at $(t, x) \in [t_0, T] \times \mathbb{R}$ given and fixed, we will identify t_0 with t , and x_0 with x , in the statically optimal control $u_*^s(t, x)$ given in (3.45) above. This yields $\tilde{H}_t = H_t/H_{t_0} = 1$, and a direct differentiation of the right-hand side in (3.40) shows that

$$(3.48) \quad \frac{\partial P}{\partial h}(T-t, 1) = -e^{r(T-t)}\mathbb{E}[H_{T-t}^2 I(\lambda > H_{T-t})]$$

for $t \in [t_0, T]$. Inserting (3.48) into (3.45), with $(t_0, x_0) = (t, x)$ and $\tilde{H}_t = 1$, upon recalling (3.3)+(3.4) and noting from (3.40) that

$$(3.49) \quad P(T-t, 1) = e^{r(T-t)}\mathbb{E}[H_{T-t}(\lambda - H_{T-t})^+],$$

we see that a candidate for the dynamically optimal control in (2.6) is given by

$$(3.50) \quad u_*^d(t, x) = \frac{\delta}{\sigma} \frac{1}{x} (x - e^{-r(T-t)}g) a(T-t, x)$$

(as stated in (3.6) above), where $a(T-t, x)$ is given in (3.7), with $\lambda = \lambda(T-t, x)$ being a unique strictly positive solution to (3.8) for $(t, x) \in [t_0, T] \times \mathbb{R}$ given and fixed (as claimed in part (A) of the theorem).

To see that the control u_*^d is dynamically optimal in the sense of Definition 2, take any control v from $C_{g, \beta}$ such that $v(t, x) \neq u_*^d(t, x)$ for $(t, x) \in [t_0, T] \times \mathbb{R}$ given and fixed, and choose w in $C_{g, \beta}$ to be the statically optimal control under the measure $\mathbb{P}_{t, x}$. Then $w(t, x) = u_*^s(t, x) = u_*^d(t, x) \neq v(t, x)$, so that by continuity of w and v we can conclude that $w(s, y) \neq v(s, y)$ for all $(s, y) \in [t, t+\varepsilon] \times [x-\varepsilon, x+\varepsilon]$ for some $\varepsilon > 0$ sufficiently small. But then by the uniqueness part of the Itô–Clark theorem, referred to following (3.23) above, we can conclude that $\mathbb{P}_{t, x}(X_T^v \neq X_T^w) > 0$, where we recall that $X_T^w = X_*$ under $\mathbb{P}_{t, x}$ with X_* given in (3.27). Moreover, since X_* is a unique point of the minimum in (3.26) up to a set of $\mathbb{P}_{t, x}$ -measure zero, it thus follows from (3.14) and (3.28) that

$$(3.51) \quad \text{Var}_{t, x}(X_T^w) = \mathbb{E}_{t, x}[(X_T^w)^2] - \beta^2 < \mathbb{E}_{t, x}[(X_T^v)^2] - \beta^2 = \text{Var}_{t, x}(X_T^v),$$

as needed in Definition 2. This shows that the control u_*^d from (3.50) is dynamically optimal in (2.6) as claimed in part (A) of the theorem.

(B) Inserting the dynamically optimal control u_*^d into (2.3) we find that the dynamically optimal wealth process X^d solves the stochastic differential equation (3.9) under the initial condition $X_{t_0}^d = x_0 \in (e^{-r(T-t_0)}g, e^{-r(T-t_0)}\beta)$. By Itô's existence and uniqueness result for locally Lipschitz coefficients (see, e.g., [19, p. 132]) we know that the stochastic differential equation (3.9) has a unique strong solution until the first exit time of X_t^d from the interval $(e^{-r(T-t)}g, e^{-r(T-t)}\beta)$ when t runs from t_0 to T followed by the constant motion corresponding to $u_*^d = 0$ afterwards (as discussed prior to Definition 1 above).

9. In order to analyze the hitting times τ_β and τ_g of X^d to the upper and lower boundaries as defined in (3.10) and (3.11), respectively, we will consider the process Z defined by

$$(3.52) \quad Z_t := e^{r(T-t)} X_t^d$$

for $t \in [t_0, T]$ (omitting superscript d from Z to simplify the notation). From (3.6)–(3.8) and (3.19) we see that Z solves

$$(3.53) \quad dZ_t = \delta(Z_t - g) A(T-t, Z_t) d\tilde{W}_t$$

under $\tilde{\mathbb{P}}_{t_0, z_0}$ for $z_0 \in (g, \beta)$ where the function A is defined by

$$(3.54) \quad A(T-t, z) := a(T-t, e^{-r(T-t)}z) = \frac{\mathbb{E}[H_{T-t}^2 I(\lambda > H_{T-t})]}{\mathbb{E}[H_{T-t}(\lambda - H_{T-t})^+]},$$

and $\lambda = \lambda(T-t, e^{-r(T-t)}z)$ is the unique strictly positive solution to (3.8), which reads

$$(3.55) \quad \frac{\mathbb{E}[H_{T-t}(\lambda - H_{T-t})^+]}{\mathbb{E}[(\lambda - H_{T-t})^+]} = e^{-r(T-t)} \frac{z - g}{\beta - g}$$

for $(t, z) \in [t_0, T] \times [g, \beta]$.

10. Note that Z is a bounded martingale and hence $\tilde{\mathbb{P}}_{t_0, z_0}$ -a.s. convergent as $t \uparrow T$ for $z_0 \in (g, \beta)$. If its limit Z_T were to belong to (g, β) , then using (3.33) we would have $\lambda(T-t, e^{-r(T-t)}Z_t) \rightarrow \lambda(0+, Z_T) \in (0, \infty)$ $\tilde{\mathbb{P}}_{t_0, z_0}$ -a.s. as $t \uparrow T$, so that the left-hand side of (3.55) would tend to 1, while its right-hand side would tend to $(Z_T - g)/(\beta - g) \in (0, 1)$, which is a contradiction. We can therefore conclude that $Z_t \rightarrow Z_T \in \{g, \beta\}$ $\tilde{\mathbb{P}}_{t_0, z_0}$ -a.s. as $t \uparrow T$.

11. We now show that Z does not hit β strictly before T with a strictly positive probability as claimed in the first equality of (3.12). For this, we first establish that the following limiting relation is satisfied:

$$(3.56) \quad \lim_{z \uparrow \beta} \frac{A(T-t, z)}{\beta - z} = \frac{1}{\beta - g} \lim_{\lambda \uparrow \infty} \left(\frac{\frac{\mathbb{E}[H_{T-t}^2 I(\lambda > H_{T-t})]}{\mathbb{E}[H_{T-t}(\lambda - H_{T-t})^+]}}{1 - e^{r(T-t)} \frac{\mathbb{E}[H_{T-t}(\lambda - H_{T-t})^+]}{\mathbb{E}[(\lambda - H_{T-t})^+]}} \right) \\ = \frac{\mathbb{E}(H_{T-t}^2)}{(\beta - g)\mathbb{E}(H_{T-t})[e^{r(T-t)}\mathbb{E}(H_{T-t}^2) - \mathbb{E}(H_{T-t})]},$$

where in the first equality we use (3.54) and the fact that $\beta - z = (\beta - g)[1 - e^{r(T-t)}F(\lambda)]$, upon letting $F(\lambda)$ denote the left-hand side in (3.55), combined with $\lambda = \lambda(T-t, e^{-r(T-t)}z) \uparrow \infty$ as $z \uparrow \beta$, which is seen from (3.55) using (3.33) above,

and the second equality follows by calculating the limit upon using the fact that the probability density function of H_{T-t} tends faster to zero than any polynomial (of third degree) when the argument tends to infinity. (Details of these lengthy but elementary calculations are omitted for simplicity.) Moreover, the convergence in (3.56) is uniform over $t \in [t_0, T_1]$ so that finiteness of the limit in (3.56) implies that the function \tilde{A} defined by

$$(3.57) \quad \tilde{A}(T-t, z) := (z-g) \frac{A(T-t, z)}{\beta-z}$$

is bounded for $(t, z) \in [t_0, T_1] \times (g, \beta)$, as it is easily seen that $(z-g)A(T-t, z) \rightarrow 0$ as $z \downarrow g$, and this convergence is uniform over $t \in [t_0, T_1]$ for any $T_1 \in [t_0, T)$ given and fixed. From (3.53) we see that Z solves

$$(3.58) \quad dZ_t = \delta(\beta - Z_t) \tilde{A}(T-t, Z_t) d\tilde{W}_t,$$

so that by setting $Y_t := \beta - Z_t$ we see that Y solves

$$(3.59) \quad dY_t = \delta Y_t \tilde{A}(T-t, \beta - Y_t) d\hat{W}_t,$$

where we set $\hat{W} = -\tilde{W}$. From (3.59) we find that

$$(3.60) \quad Y_t = Y_{t_0} \exp \left[\int_{t_0}^t \tilde{A}(T-s, \beta - Y_s) d\hat{W}_s - \frac{1}{2} \int_{t_0}^t \tilde{A}^2(T-s, \beta - Y_s) ds \right],$$

with $\int_{t_0}^t \tilde{A}^2(T-s, \beta - Y_s) ds < \infty$ for $t \in [t_0, T_1]$ by the boundedness of \tilde{A} . Hence from (3.60) we see that $Y_t > 0$, i.e., $Z_t < \beta$ \hat{P}_{t_0, z_0} -a.s. for all $t \in [t_0, T_1]$ with $T_1 \in [t_0, T)$ and $z_0 \in (g, \beta)$, showing that Z does not hit β strictly before T with a strictly positive probability, as claimed in the first equality in (3.12). The second equality in (3.12) follows from the fact that Z is a bounded martingale converging \hat{P}_{t_0, z_0} -a.s. to either g or β as $t \uparrow T$ for $z_0 \in (g, \beta)$, so that Z must hit β at time T with a strictly positive probability as claimed.

12. We finally show that Z hits g strictly before T with a strictly positive probability as claimed in (3.13). For this, we first establish that the following limiting relation is satisfied:

$$(3.61) \quad \lim_{z \downarrow g} \frac{A(T-t, z)}{|\log(z-g)|} = \lim_{\lambda \downarrow 0} \frac{\mathbf{E}[H_{T-t}^2 I(\lambda > H_{T-t})]}{|\log(\lambda)| \mathbf{E}[H_{T-t}(\lambda - H_{T-t})^+]} = 1,$$

where in the first equality we use that $\log(z-g) = \log(\beta-g) + r(T-t) + \log(F(\lambda))$, upon letting $F(\lambda)$ denote the left-hand side in (3.55), combined with $\lambda = \lambda(T-t, e^{-r(T-t)}z) \downarrow 0$ as $z \downarrow g$, which is seen from (3.55) using (3.33) above, and that

$$(3.62) \quad \lim_{\lambda \downarrow 0} \frac{F(\lambda)}{\lambda} = \lim_{\lambda \downarrow 0} \frac{\mathbf{E}[H_{T-t}(\lambda - H_{T-t})^+]}{\lambda \mathbf{E}[(\lambda - H_{T-t})^+]} = 1,$$

upon using the fact that $\log(H_{T-t})$ is normally distributed to establish the final equalities in (3.61) and (3.62). (Details of these lengthy but elementary calculations are omitted for simplicity.) Moreover, the convergence in (3.61) is uniform over $t \in [t_0, T_1]$, so that the strict positivity and finiteness of the limit in (3.61) imply that the function \tilde{A} defined by

$$(3.63) \quad \tilde{A}(T-t, z) := \frac{A(T-t, z)}{|\log(z-g)|}$$

is bounded away from zero and bounded from above for $(t, z) \in [t_0, T_1] \times (g, \beta)$ with any $T_1 \in [t_0, T)$ given and fixed. From (3.53) we see that Z solves

$$(3.64) \quad dZ_t = \delta(Z_t - g) |\log(Z_t - g)| \tilde{A}(T - t, Z_t) d\tilde{W}_t,$$

so that by setting $Y_t := Z_t - g$, we see that Y solves

$$(3.65) \quad dY_t = \delta Y_t |\log Y_t| \tilde{A}(T - t, g + Y_t) d\tilde{W}_t.$$

Defining the additive functional

$$(3.66) \quad C_t := \int_{t_0}^t \tilde{A}^2(T - s, g + Y_s) ds$$

and letting $S_t := C_t^{-1}$, we find that the time-changed process $\hat{Y}_t := Y_{S_t}$ solves

$$(3.67) \quad d\hat{Y}_t = \delta \hat{Y}_t |\log \hat{Y}_t| d\hat{W}_t,$$

where $\hat{W}_t = \int_{t_0}^t \tilde{A}(T - S_s, g + Y_{S_s}) d\tilde{W}_{S_s}$ defines a standard Brownian motion for $t \in [t_0, T]$, as is easily seen by Lévy’s characterization theorem. The “logarithmic” diffusion process \hat{Y} solving (3.67) is on natural scale (local martingale), so that the scale function of \hat{Y} is given by $S(y) = y$ for $y \in (0, \infty)$. Since $S(\infty) = \infty$, we see that the explosion of \hat{Y} (i.e. its exit from $(0, \infty)$) in finite time is only possible at 0. Moreover, we see that

$$(3.68) \quad v(y) := \int_1^y S'(z) \left(\int_1^z \frac{2 dw}{S'(w) \sigma^2(w)} \right) dz = \int_1^y \int_1^z \frac{2 dw dz}{\delta^2 w^2 \log^2 w}$$

for $y \in (0, \infty)$, where $\sigma(w) := \sigma w |\log w|$ for $w \in (0, \infty)$. A direct examination of the final (double) integral in (3.68) shows that

$$(3.69) \quad v(0+) < \infty \quad \text{and} \quad v(\infty) = \infty.$$

Therefore by Feller’s test (see [9, Prop. 5.32, p. 350]) we can conclude that the explosion of \hat{Y} at zero occurs with probability one (i.e., \hat{Y} hits 0 with probability one). Since \hat{Y} is a regular time-homogeneous diffusion process, it follows therefore that \hat{Y} hits zero with a strictly positive \hat{P}_{t_0, y_0} -probability before any given deterministic time $t_1 \in (0, T - t_0]$ has elapsed. Using the fact that $\tilde{A}(T - t, z)$ is bounded away from zero for $(t, z) \in [t_0, T_1] \times (g, \beta)$ with any $T_1 \in [t_0, T)$ given and fixed, from (3.66) we see that $C_t \geq c(t - t_0) =: t_1$ for $t \in (t_0, T_1]$ with some $c > 0$. This shows that C_t is greater than or equal to the first hitting time of \hat{Y} to 0 with a strictly positive probability. Since 0 is an exit (not an entrance) boundary point for \hat{Y} , this shows that $\hat{Y}_{t_0 + C_t} = Y_t = 0$ with a strictly positive probability for $t \in (t_0, T_1]$. Thus Z hits g strictly before T with a strictly positive probability as claimed in (3.13), and the proof of the theorem is complete. \square

Remark 4. The proof above shows that there exists a *unique* pair (λ_1, λ_2) in \mathbb{R}^2 such that X_* from (3.27) solves (3.24)+(3.25), where λ_2 is always strictly negative (while λ_1 is real valued). For this reason, following (3.27) we have imposed that $\kappa > 0$, i.e., $\lambda_2 < 0$ and sought the solution (λ_1, λ_2) to (3.28)+(3.29) in $\mathbb{R} \times (-\infty, 0)$. The arguments presented in parts 3 and 4 of the proof above establish the existence of a unique pair (λ_1, λ_2) in $\mathbb{R} \times (-\infty, 0)$ such that X_* solves (3.24)+(3.25). To

complete the proof of the initial claim above, it is enough to show that there is no $\lambda_2 \geq 0$ and $\lambda_1 \in \mathbb{R}$ such that (3.28) and (3.29) hold with $+H$ in place of $-H$, where $\kappa = \lambda_2/2$ and $\lambda = (\lambda_1/2 - g)/(\lambda_2/2)$. This verification can be made along the same lines as the proof above. Indeed, defining F as in (3.32) with $+H$ in place of $-H$, it is easily seen using the same arguments as above that $\lambda \mapsto F(\lambda)$ is (strictly) decreasing on $(-\infty, 0]$. Denoting the right-hand side in (3.31) by R , and noting that $R \in (0, E(H))$, we thus see that if the equality in (3.31) were to hold for some $\lambda \leq 0$, then we would have $E(H^2)/E(H) = F(0) \leq F(\lambda) = R < E(H)$, which would violate Jensen's inequality. Similarly, if equality in (3.31) were to hold for some $\lambda > 0$, then by solving for λ in (3.31) we would get $\lambda = -(E(H^2) - E(H)R)/(E(H) - R) > 0$, which would again violate Jensen's inequality due to $R < E(H)$. Combined with the arguments recalled above this shows that there exists a unique pair (λ_1, λ_2) in \mathbb{R}^2 such that X_* solves (3.24)+(3.25) as claimed.

Moreover, it may be noted that X_* is almost surely a *pointwise unique* solution to (3.24)+(3.25). Indeed, if Z_* solves (3.24)+(3.25) then denoting the value in (3.24) by V , we see that $V = E(Z_*^2) = E(Z_*^2 - \lambda_1 Z_* - \lambda_2 H Z_*) + \lambda_1 \beta + \lambda_2 x \geq E(X_*^2 - \lambda_1 X_* - \lambda_2 H X_*) + \lambda_1 \beta + \lambda_2 x = E(X_*^2) = V$, from which we infer that if the inequality $Z_*^2 - \lambda_1 Z_* - \lambda_2 H Z_* \geq X_*^2 - \lambda_1 X_* - \lambda_2 H X_*$ were to be strict with a strictly positive probability, then we would have $V > V$, which is a contradiction. Hence $Z_* = X_*$ almost surely by the uniqueness of the minimum point X_* of $x \mapsto x^2 - \lambda_1 x - \lambda_2 H x$ for $x \geq g$ as claimed.

Remark 5. Note that one can extract the statically optimal control u_*^s , the statically optimal controlled wealth process X^s , and the static value function V_s from the proof above. They are given by (3.45), (3.41), and (3.47), respectively (compare these expressions with the pathwise unconstrained ones given in [16, Corollary 7, Part A]). Note, in particular, that in addition to depending explicitly on the initial time t_0 and the initial wealth x_0 , the statically optimal control u_*^s is expressed explicitly in terms of the statically optimal controlled wealth process X_t^s evaluated at the current time t (and in terms of the current time t itself).

The dynamically optimal control u_*^d from (3.6) by its nature rejects any past point (t_0, x_0) to measure its performance, so that any comparison between the static value $V_s(t_0, x_0)$ and the dynamic value $V_d(t_0, x_0)$ is meaningless from the standpoint of the dynamic optimality. Another issue with a plain comparison between the values $V_s(t, x)$ and $V_d(t, x)$ for $(t, x) \in [t_0, T] \times \mathbb{R}$ is that the optimally controlled processes X^s and X^d may never arrive at the same point x at the same time t , so that the comparison itself may be unreal. A more dynamic method, which also makes more sense in general, is to compare the value functions composed with the controlled processes. This amounts to looking at $V_s(t, X_t^s)$ and $V_d(t, X_t^d)$ for $t \in [t_0, T]$ and paying particular attention to t becoming the terminal value T .

Recalling that X_T^d takes only values g and β , and noting that $V_d(t, g) = V_d(t, \beta) = 0$ for all $t \in [t_0, T]$, we see that $V_d(T, X_T^d) = 0$. On the other hand, it is easily seen from (3.31) with $H = H_T/H_t$ that $\lambda = \lambda(T - t, X_t^s) \rightarrow 0$ as $t \uparrow T$ when $X_T^s \in (g, \beta)$. From (3.27)+(3.28) and (3.46)+(3.47), we therefore find, using the first equality in (3.30) and L'Hospital's rule, that

$$(3.70) \quad \begin{aligned} V_s(t, X_t^s) &= \text{Var}_{t, X_t^s}(X_T^s) = E[(\kappa(\lambda - H)^+ + g)^2] - \beta^2 \\ &= (\beta - g)^2 \left[\frac{E[(\lambda - H)^+]^2}{(E[(\lambda - H)^+])^2} - 1 \right] \rightarrow \infty \end{aligned}$$

as $t \uparrow T$ whenever $X_T^s \in (g, \beta)$ due to $\lambda = \lambda(T - t, X_t^s) \rightarrow 0$ and $H = H_T/H_t \rightarrow 1$.

However, from (3.46) we see that $X_T^s \in (g, \beta)$ with a strictly positive \mathbb{P}_{t_0, x_0} -probability, so that $\mathbb{E}_{t_0, x_0}[V_s(t, X_t^s)] \rightarrow \infty$ as $t \uparrow T$. It follows therefore that

$$(3.71) \quad E_{t_0, x_0}[V_s(T, X_T^s)] = \infty > 0 = E_{t_0, x_0}[V_d(T, X_T^d)]$$

for $(t_0, x_0) \in [0, T] \times \mathbb{R}$ such that $e^{-r(T-t_0)}g < x_0 < e^{-r(T-t_0)}\beta$. This shows that the dynamic control u_*^d from (3.6) outperforms the static control u_*^s from (3.45) in the pathwise constrained problem (2.6).

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